

# Ordinal response regression models

Categorical Data Analysis, Summer 2015

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## Overview

- In the previous lecture, we discussed the baseline-category logit model for multivariable regression modeling of a multinomial response variable  $Y$  with  $J$  categories.
- Recall that this model estimates a complete set of  $p$  parameters (intercept and slopes) for each of  $J - 1$  baseline-referenced logits.
- When the response variable is ordinal, we can use this structure to our advantage to develop more parsimonious and powerful models.

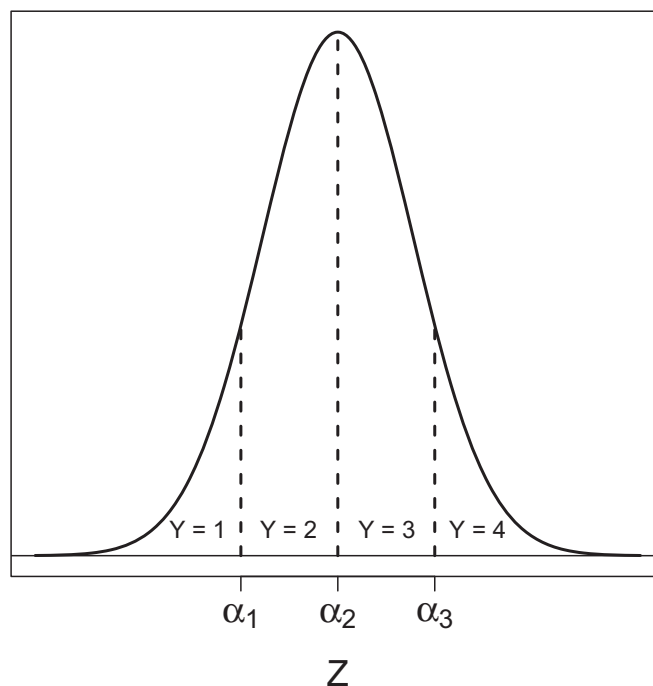
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## Ordinal variables

- Ordinal outcomes are common in medical research
  - Likert scale ('Strongly disagree', 'Disagree', 'Neither agree nor disagree', 'Agree', 'Strongly agree')
  - Disease severity ('Normal', 'Mild', 'Moderate', 'Severe')
- It is often possible to consider an ordinal variable to be a discretized version of a continuous **latent variable**.
- In this context, the ordinal variable is a discrete version of an unmeasured (and unobserved) continuous variable.

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## Ordinal variables (cont.)



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## Cumulative logits

- A natural approach for regression modeling with an ordinal response variable is to construct logit models for dichotomized versions of  $Y$  while moving in sequence 'up' the ordinal scale.
- Such logits are called **cumulative logits** and are constructed from cumulative probabilities

$$P(Y \leq j|\mathbf{x}) = \pi_1(\mathbf{x}) + \dots + \pi_j(\mathbf{x}), j = 1, \dots, J$$

- The cumulative logits are defined as

$$\begin{aligned} \text{logit}[P(Y \leq j|\mathbf{x})] &= \log \frac{P(Y \leq j|\mathbf{x})}{P(Y > j|\mathbf{x})} \\ &= \log \frac{P(Y \leq j|\mathbf{x})}{1 - P(Y \leq j|\mathbf{x})} \end{aligned}$$

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## Cumulative logits (cont.)

In contrast to the baseline-category logits, each cumulative logit uses all response categories.

$$\begin{aligned} L_1 &= \text{logit}[P(Y \leq 1|\mathbf{x})] = \log \frac{\pi_1(\mathbf{x})}{\pi_2(\mathbf{x}) + \pi_3(\mathbf{x}) + \dots + \pi_J(\mathbf{x})} \\ L_2 &= \text{logit}[P(Y \leq 2|\mathbf{x})] = \log \frac{\pi_1(\mathbf{x}) + \pi_2(\mathbf{x})}{\pi_3(\mathbf{x}) + \dots + \pi_{J-1}(\mathbf{x}) + \pi_J(\mathbf{x})} \\ &\quad \vdots \\ L_{J-1} &= \text{logit}[P(Y \leq J-1|\mathbf{x})] = \log \frac{\pi_1(\mathbf{x}) + \pi_2(\mathbf{x}) + \dots + \pi_{J-1}(\mathbf{x})}{\pi_J(\mathbf{x})} \end{aligned}$$

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## The Proportional Odds Model

- Modeling a single cumulative logit as a function of covariates is equivalent to logistic regression where categories  $1, \dots, j$  and  $j + 1, \dots, J$  form the outcomes of success and failure, respectively.
- The **proportional odds model** simultaneously uses all  $J - 1$  cumulative logits in a single model, and is given by

$$\text{logit}[P(Y \leq j|\mathbf{x})] = \alpha_j + \beta'\mathbf{x}, \quad j = 1, \dots, J - 1$$

- The natural ordering of the cumulative probabilities results in model constraints:

$$\begin{aligned} P(Y \leq 1|\mathbf{x}) &\leq \dots \leq P(Y \leq J - 1|\mathbf{x}) \\ \text{logit}[P(Y \leq 1|\mathbf{x})] &\leq \dots \leq \text{logit}[P(Y \leq J - 1|\mathbf{x})] \\ \alpha_1 + \beta'\mathbf{x} &\leq \dots \leq \alpha_{J-1} + \beta'\mathbf{x} \end{aligned}$$

- We therefore constrain the  $\{\alpha_j\}$  so that  $\alpha_1 \leq \dots \leq \alpha_{J-1}$ .

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## The Proportional Odds Model (cont.)

- Additionally, the proportional odds model assumes the same effects  $\beta$  across all cumulative logits.
- The common  $\beta$  and the ordering constraint on the intercepts are important differences between the proportional odds model and the baseline-category logit model.
- Cumulative probabilities are constructed using the inverse logit function

$$P(Y \leq j|\mathbf{x}) = \frac{\exp\{\alpha_j + \beta'\mathbf{x}\}}{1 + \exp\{\alpha_j + \beta'\mathbf{x}\}}, \quad j = 1, \dots, J - 1$$

- Individual category probabilities are constructed by taking differences of cumulative probabilities

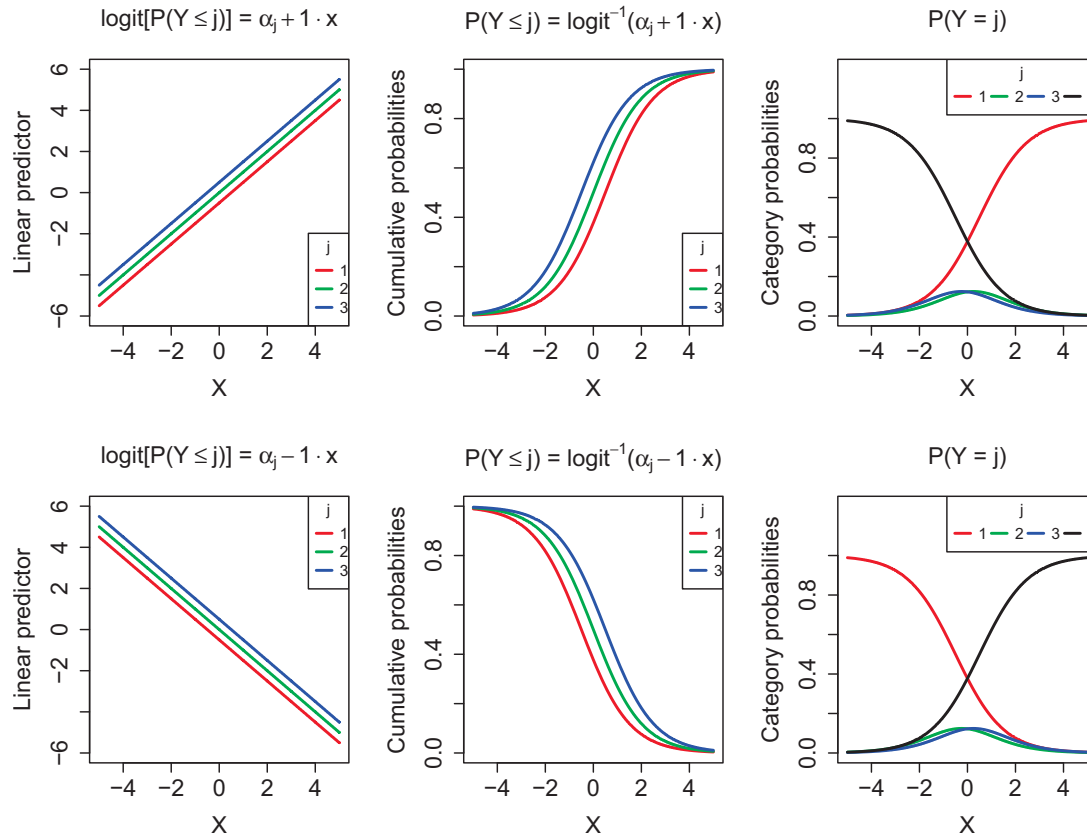
$$P(Y = 1|\mathbf{x}) = P(Y \leq 1|\mathbf{x})$$

$$P(Y = j|\mathbf{x}) = P(Y \leq j|\mathbf{x}) - P(Y \leq j - 1|\mathbf{x}), \quad j = 2, \dots, J - 1$$

$$P(Y = J|\mathbf{x}) = 1 - P(Y \leq J - 1|\mathbf{x})$$

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## Four-category proportional odds model



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## Observations

- The proportional odds assumption implies parallel linear predictors for the  $J - 1$  cumulative logits (parallelism is not a property of baseline-category logits).
- Changing the sign of  $\beta$  results in a reverse ordering of the response categories.
- This means that the model 'preserves' the ordinal scale of  $Y$ , and is therefore an attractive model for regression models of ordinal response variables.

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## Parameter interpretation

Consider a single continuous predictor  $x$ . The difference in cumulative logits for an increment  $\Delta$  of  $x$  is given by

$$\begin{aligned}\text{logit}([P(Y \leq j|x = x_0 + \Delta)]) - \text{logit}([P(Y \leq j|x = x_0)]) &= \\ (\alpha_j + \beta(x_0 + \Delta)) - (\alpha_j + \beta x_0) &= \Delta\beta\end{aligned}$$

- $e^{\Delta\beta}$  is a **cumulative odds ratio**
- The cumulative odds ratio does not depend on  $j$ , and is therefore the same across all  $J - 1$  cumulative logits.
- This means that *no matter how you dichotomize  $Y$* , the effect of  $x$  on the odds that  $Y \leq j$  is constant.
- *This property of a common effect for all cumulative probabilities is referred to as **proportional odds**.*

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## Parameter interpretation (cont.)

- Since cumulative odds ratios in proportional odds models are the same for all categories, subsequent inference focuses on the *direction* of response rather than on specific response categories.
- Suppose  $e^{\Delta\beta} = 2$  for  $\Delta = 1$ .
  - This cumulative odds ratio is interpreted as a two-fold increase in the odds of a response as small or smaller for a 1-unit increase in  $x$ .
  - Alternatively, since  $1/e^{\Delta\beta} = 0.5$ , we could also conclude that there is a 50% reduction in the odds of a higher response for a 1-unit increase in  $x$ .
- Suppose  $e^{\Delta\beta} = 0.6$  for  $\Delta = 1$ .
  - There is a 40% reduction in the odds of a response as small or smaller for a 1-unit increase in  $x$ .
  - Alternatively, since  $1/e^{\Delta\beta} \doteq 1.67$ , we could also conclude that there is a 67% increase in the odds of a higher response for a 1-unit increase in  $x$ .

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