(MODIFIED) TOXICITY PROBABILITY INTERVALS JI ET AL., Clinical Trials (2007, 2010)

Jody Ciolino

Division of Biostatistics and Epidemiology Medical University of South Carolina

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THE 3+3 DESIGN STILL PERSISTS

- Between April and May of 2006, 20/22 phase I clinical trials submitted to IRB (at MD Andersons) called for 3+3 design!
- Why?
 - SIMPLE: Investigators can design trials without help of statistician
 - Large volume of phase I trials \Rightarrow impossible to provide every phase I clinical trial with statistician
 - Computer required throughout for sophisticated designs like CRM
 - In CRM design, simulations required before for calibration

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- Develop a method with better performance than 3+3 design
- BUT easily understood and implemented
- Known as (modified) Toxicity Probability Interval (mTPI)

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THE MODEL

- Let p_T = the target toxicity probability for a drug
- Goal = find highest dose with toxicity closest to p_T (assume monotonic relationship between dose and toxicity as well as between dose and efficacy)

• Let
$$p_i = p(toxicity)$$
 at dose $i = 1, ..., d$

•
$$l(p) \propto \prod_{i=1}^d p_i^{x_i} (1-p_i)^{n_i-x_i}$$

- Where x_i = total number of subjects treated at dose i that experience toxicity and n_i = total number of subjects treated at dose i
- Assume that priors of p_i are i.i.d beta (α, β)

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BETA PRIORS

• Beta density:
$$f(p) = \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)}p^{\alpha-1}(1-p)^{\beta-1}$$

•
$$p \in (0, 1)$$
, mean $= \frac{\alpha}{\alpha + \beta}$, var $= \frac{\alpha \beta}{(\alpha + \beta + 1)(\alpha + \beta)^2}$

 In the Bayesian framework, the Beta prior is the conjugate prior for the binomial likelihood (posterior will also be a beta distribution)

• Posterior =
$$beta(\alpha + x_i, \beta + n_i - x_i)$$

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More on the Beta Distribution



 $\alpha = \beta = 1 \Rightarrow U(0, 1)$ Authors recommend $\alpha = \beta = 0.005$ ("non-informative").

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If patients are treated at dose i, only have 3 possible actions...

- De-escalate (D) to dose i-1
- Stay (S) at dose i
- Scalate (E) to dose i + 1

Partition the interval (0,1) into three parts, such that a posterior probability falling in each interval is too low, close to p_T , or too high, respectively:

- TPI: {(0, $p_T k_1 \sigma_i$), ($p_T k_1 \sigma_i$, $p_T + k_2 \sigma_i$), ($p_T + k_2 \sigma_i$, 1)}
- mTPI: { $(0, p_T \epsilon_1), (p_T \epsilon_1, p_T + \epsilon_2), (p_T + \epsilon_2, 1)$ }
- Choose E,D, or S depending on which interval has highest posterior mass
- It may be difficult to determine k₁ and k₂ and results may be sensitive to k₁ and k₂
- Focus on mTPI (2010) that is based on the equivalence interval $[p_{T}-\epsilon_{1},p_{T}+\epsilon_{2}]$

- Choose $\xi \in (0, 1)$ that is relatively large such that...
- If $p(p_1 > p_T | data) > \xi \Rightarrow$ terminate trial
- If decision is E and $p(p_{i+1} > p_T | data) > \xi \Rightarrow$ Do no escalate.

ESTIMATING THE MTD

- At end of trial, select dose with smallest difference $|\hat{p}_i p_T|$ among all tried doses i = 1, ..., d
- Must satisfy safety criterion: $p(p_i > p_T | data) \le \xi$
- Here, p̂_i is a "sensible" estimate of p_i...isotonically transformed posterior mean
- Ji et al (2007) discuss isotonic regression that borrows strength across doses: compute posterior mean \hat{p}_i under beta posterior distribution and then perform pooled adjacent violators algorithm (PAVA) on \hat{p}_i (Goodman et al, *Stat Med*, 1995)
- Dealing with ties: (1)Choose the highest dose that is less results in \hat{p}_i less than p_T or (2)the lowest dose tat results in \hat{p}_i greater than p_T

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MORE ON MTPI

 Define Unit Probability Mass (UPM) = the ratio of probability of interval and length of interval

• i.e. if
$$X \sim F$$
, UPM of $(a, b] = \frac{F(b) - F(a)}{b - a}$

- Calculate the UPM for each interval $\{(0, p_T \epsilon_1), (p_T \epsilon_1, p_T + \epsilon_2), (p_T + \epsilon_2, 1)\}$
- Choose E,D, or S corresponding to the interval with largest UPM
- It can be shown that this dose assignment rule minimizes the posterior expected penalty in the Bayesian framework...

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PROBABILITY MODEL: MTPI



Define penalties for each decision as follows...

$$L(D, p_i) = \begin{cases} K_D, & \text{if } -\epsilon_1 \le p_i - p_T \le \epsilon_2; \\ 0, & \text{if } p_i - p_T > \epsilon_2; \\ N_D, & \text{if } p_i - p_T < -\epsilon_1; \end{cases} \qquad L(S, p_i) = \begin{cases} 0, & \text{if } -\epsilon_1 \le p_i - p_T \le \epsilon_2; \\ M_S, & \text{if } p_i - p_T > \epsilon_2; \\ N_S, & \text{if } p_i - p_T < -\epsilon_1; \end{cases}$$
$$L(E, p_i) = \begin{cases} K_E, & \text{if } -\epsilon_1 \le p_i - p_T \le \epsilon_2; \\ M_E, & \text{if } p_i - p_T > \epsilon_2; \\ 0, & \text{if } p_i - p_T < -\epsilon_1. \end{cases}$$

PROBABILITY MODEL: MTPI

- $K_D, K_E, M_S, M_E, N_S, N_B$ are all positive real numbers
- Let X = {(x₁, n₁), ..., (x_d, n_d)} be accumulated data and let the information set corresponding to X be F
- Define R(D, p_i) = E{L(D, p_i)|F}, R(S, p_i) = E{L(S, p_i)|F}, R(E, p_i) = E{L(E, p_i)|F} as the corresponding posterior expected penalies
- Let $q_{D_i} = p(p_i p_T > \epsilon_2 | F)$, $q_{S_i} = p(-\epsilon_1 \le p_i p_T \le \epsilon_2 | F)$, $q_{E_i} = p(p_i - p_T < -\epsilon_1 | F)$
- Then $R(D, p_i) = K_D q_{S_i} + N_D q_{E_i}$, $R(S, p_i) = M_S q_{D_i} + N_S q_{E_i}$, $R(E, p_i) = K_E q_{S_i} + M_E q_{D_i}$
- We want to minimize the expected penalty, and thus we choose the action that corresponds to minimizing $R(m, p_i)$, where $m \in D, S, E$

PROBABILITY MODEL: MTPI

• Set
$$K_D = K_E = \frac{1}{\epsilon_1 + \epsilon_2}$$
, $M_S = M_E = \frac{1}{1 - p_T - E_2}$,
 $N_D = N_S = \frac{1}{p_T - \epsilon_1}$

- Under these conditions, prior expected penalties for D, E, and S are the same (one action is not favored over another a priori)
- Also, under these conditions, the posterior expected penalties $R(D, p_i)$, $R(S, p_i)$, and $R(E, p_i) = 1$ -UPMs for intervals $(0, p_T \epsilon_1)$, $[p_T \epsilon_1, p_T + \epsilon_2]$, $(p_T + \epsilon_2, 1)$
- Thus, the interval with the largest UPM will determine the decision to E, D, or S
- Large sample properties: mTPI will choose correct dose in large samples and when enough patients have been treated, this design will always choose a dose in the equivalence interval to treat all future patients (given that this dose is one of target candidates in the trial)

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COMPARING TPI, MTPI, AND CRM

- Simulations in Ji et al. suggest comparable performance between CRM and TPI, but 3+3 design="worst"
- Criteria: Percent of trials choosing correct MTD, toxicity percentage
- Several scenarios examined
- In Ji et al. (2010), mTPI showed lowest toxicity percentage in all but 1 scenario (for model based scenarios)
- 3+3 design = conservative (low toxicity percentage)

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- Robust to different values of ϵ_1 and ϵ_2 (Table 2)
- Robust to different beta priors (Table 3)
- Additional research (not shown) also suggests these results
- Independence of priors for p_i???
- Most likely dependent (since toxicity probabilities are most likely ordered), but for small sample sizes, authors believe dependence will have large influence on operating characteristics

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Software is available online at http://odin.mdacc.tmc.edu/ ylji/

Excel Macro and R programs, BUT authors cannot guarantee bug free! Be careful and check for bugs.

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CONCLUDING REMARKS

- CRM = excellent method for dose finding and has capacity to outperform (m)TPI
- BUT simplicity = key!
- mTPI over TPI...less confusing, easier parameters to define a priori
- mTPI (or TPI) may be a nice compromise between 3+3 and CRM
- Limitations: Assuming dichotomous outcome, monotonic relationship between toxicity and dose, efficacy not taken into account here
- Priors and independence of p_is

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