

# A new model for joint analysis of an event process and associated longitudinal outcome

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# Motivation

A recurrent event process has an associated longitudinal outcome.

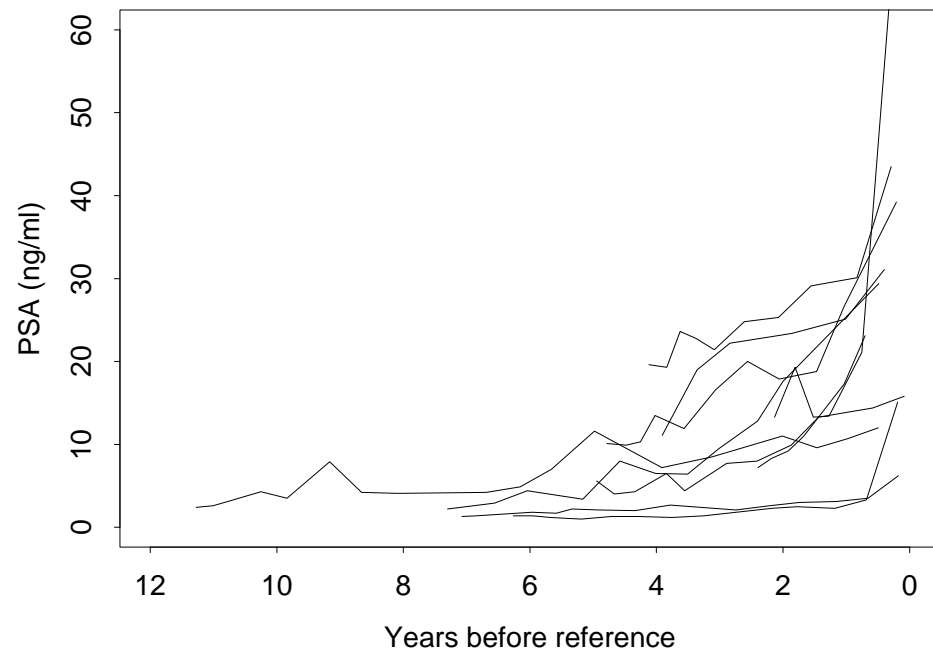
- Longitudinal outcome (marker)
  - a series of measurements for each individual
  - need to accommodate serial correlation, irregular observation times
- Recurrent event outcome
  - each individual may experience multiple events
  - need to accommodate censoring, correlation among inter-event times

# Example Scenarios

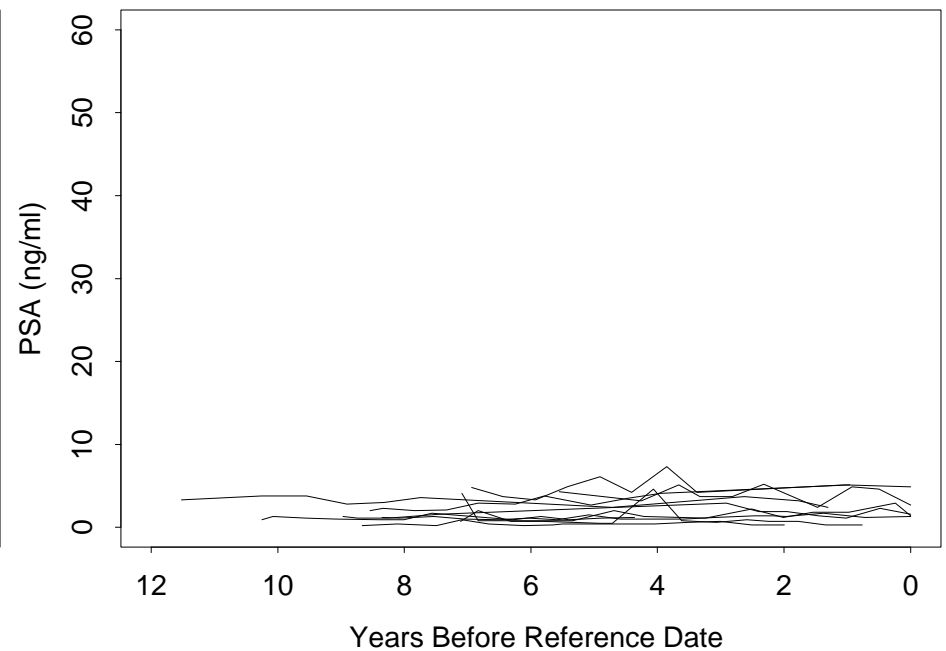
Marker	Event
PSA	PCa (single event)
Cholesterol level	MI
LVH	Heart failure
Angina	MI
Stress level	Arrhythmia
Air pollution	Asthma hospitalizations
Plasma level of new drug	disease-related events
“Chatter”	Terrorist attacks

# Examples – NPCT PSA Data

## Select PCa Cases



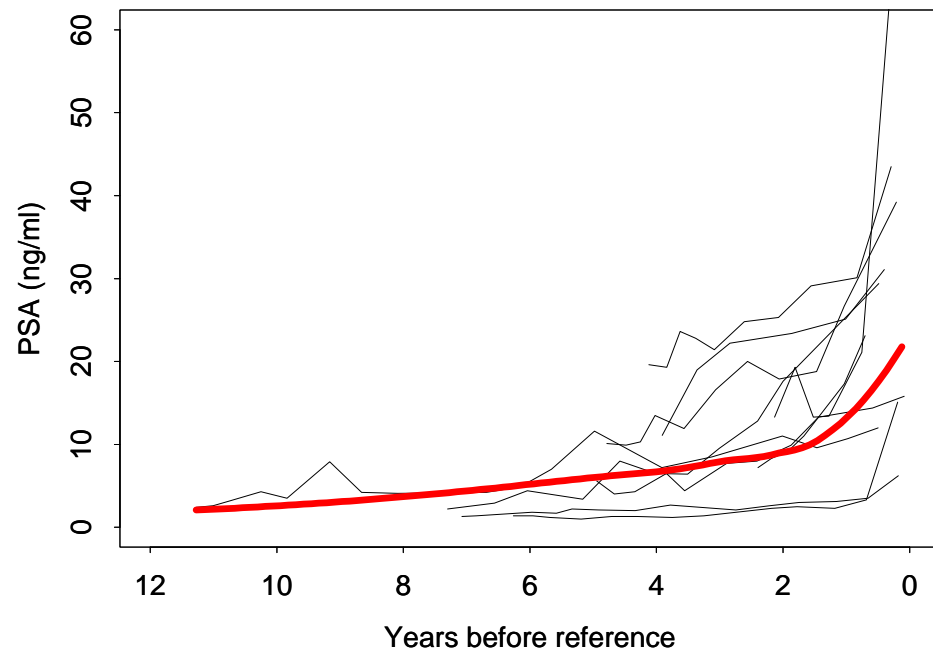
## Select Non-PCa Cases



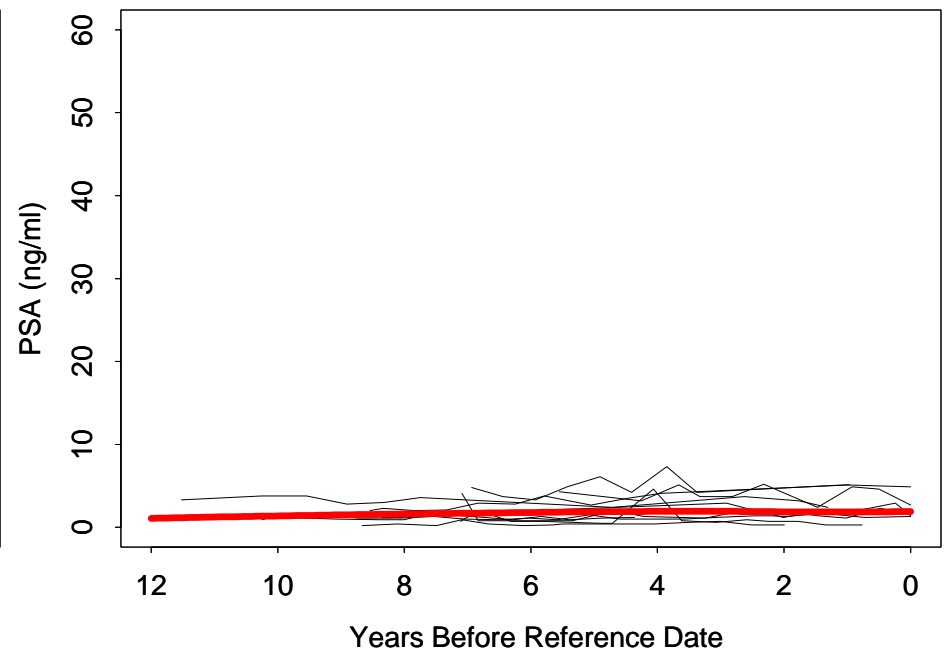
Clark, *et al.*, *JAMA*, 1996; Slate and Clark, 1999.

# Examples – NPCT PSA Data

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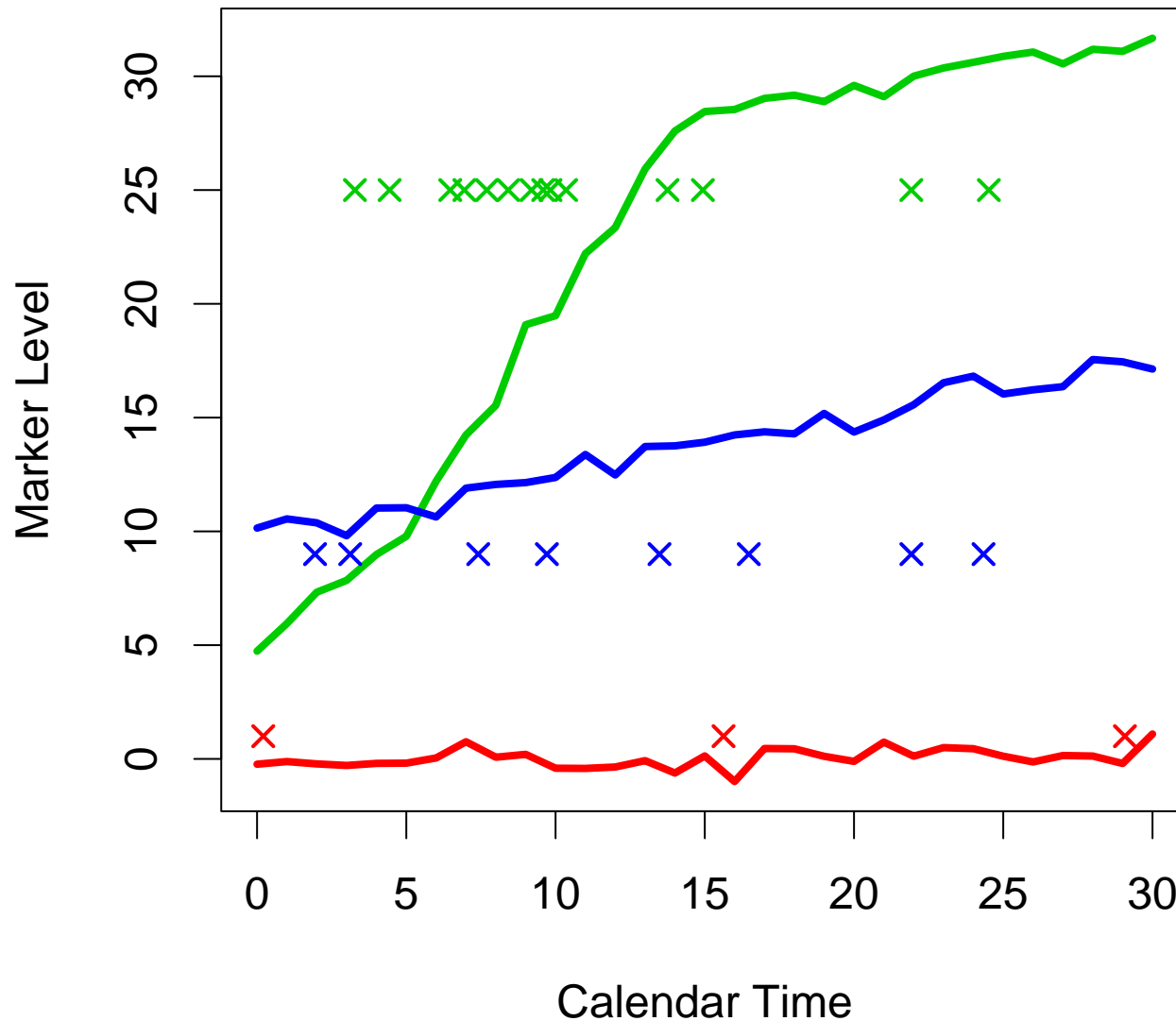


Select Non-PCa Cases



Clark, *et al.*, *JAMA*, 1996; Slate and Clark, 1999.

# Schematic – Recurrent Event



# Inferential Goals

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- What is the effect of the event history?
- What is the effect of action following the occurrence of an event?
- Does the marker aid in prediction of events?
- Is the occurrence of an event predictive for the marker?
- Is the association between the marker and event process the same for all individuals?

# Modeling Approaches

- Event Process | Marker
  - Marker as (time-dependent) covariate for event process.
  - Predict events given marker history.
- Marker | Event Process
  - Event occurrence times as covariate for marker process.
  - Characterize marker evolution for different event patterns.
- **Focus here:** (Marker, Event Process)  
modeling event process and longitudinal marker as *joint* outcomes.

# Benefits of modeling jointly

- Fitted model provides a concise description of the data available.
- Context for evaluating the association between the marker and event process.
- Model provides a means for incorporating the information from the longitudinal marker for predicting the event process; and vice versa.
- Given multiple markers, can aid selection of marker that is most informative (predictive) about the event process.

# Approaches to joint modeling

- [Event | Marker] [Marker] (measurement error model)  
e.g., CD4 counts and progression to AIDS
  - Tsiatis, DeGruttola and Wulfsohn (1995)
  - Faucett and Thomas (1996)
  - Shi, Taylor and Muñoz (1996)
  - Wulfsohn and Tsiatis (1997)
  - Law, Taylor and Sandler (2002)

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- [Event | Marker] [Marker] (measurement error model)
- [Marker | Event] [Event]  
e.g., Pawitan and Self (1993) in AIDS context
- Latent Variables to simplify dependence
  - Shared random effects  
e.g., DeGruttola and Tu (1994), Schluchter (1992)
  - Latent processes  
e.g., Xu and Zeger (2001), Henderson, Diggle and Dobson (2000)
  - Latent classes  
Lin *et al.* (2002)

# Latent variable joint modeling

Marker  $\mathbf{Y}_i(t)$  observed at times  $t_{i1}, t_{i2}, \dots, t_{im_i}$

Event Time  $T_i$  single event

- Shared random effects

$$[\mathbf{Y}_i, T_i] = \int [\mathbf{Y}_i | \mathbf{b}_i][T_i | \mathbf{b}_i][\mathbf{b}_i] d\mathbf{b}_i$$

- Latent processes

$$[\mathbf{Y}_i(t), T_i] = \int [\mathbf{Y}_i(t) | \eta_i(t)][T_i | \eta_i(t)][\eta_i(t)] d\eta_i(t)$$

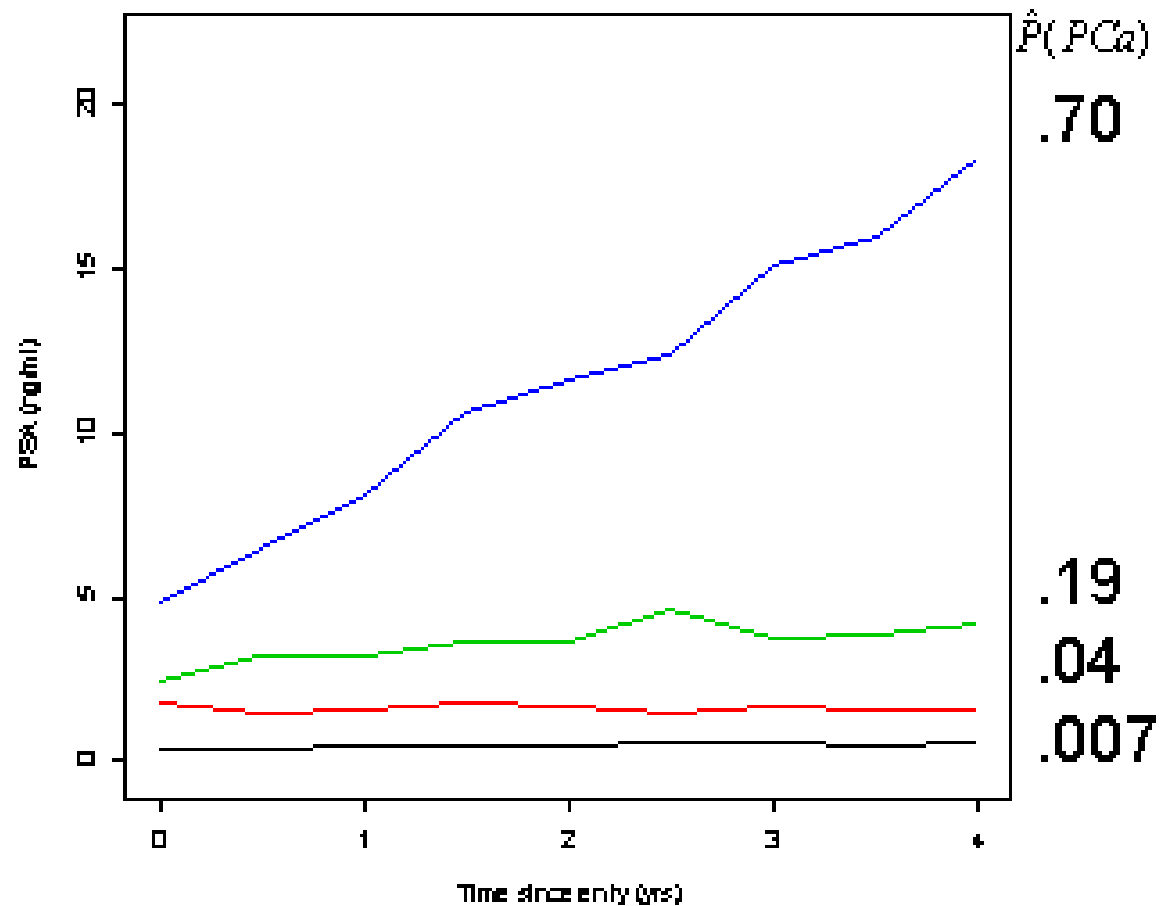
- Latent classes

$$[\mathbf{Y}_i, T_i] = \sum_{k=1}^K [\mathbf{Y}_i | \mathbf{c}_{ik}][T_i | \mathbf{c}_{ik}][\mathbf{c}_{ik}]$$

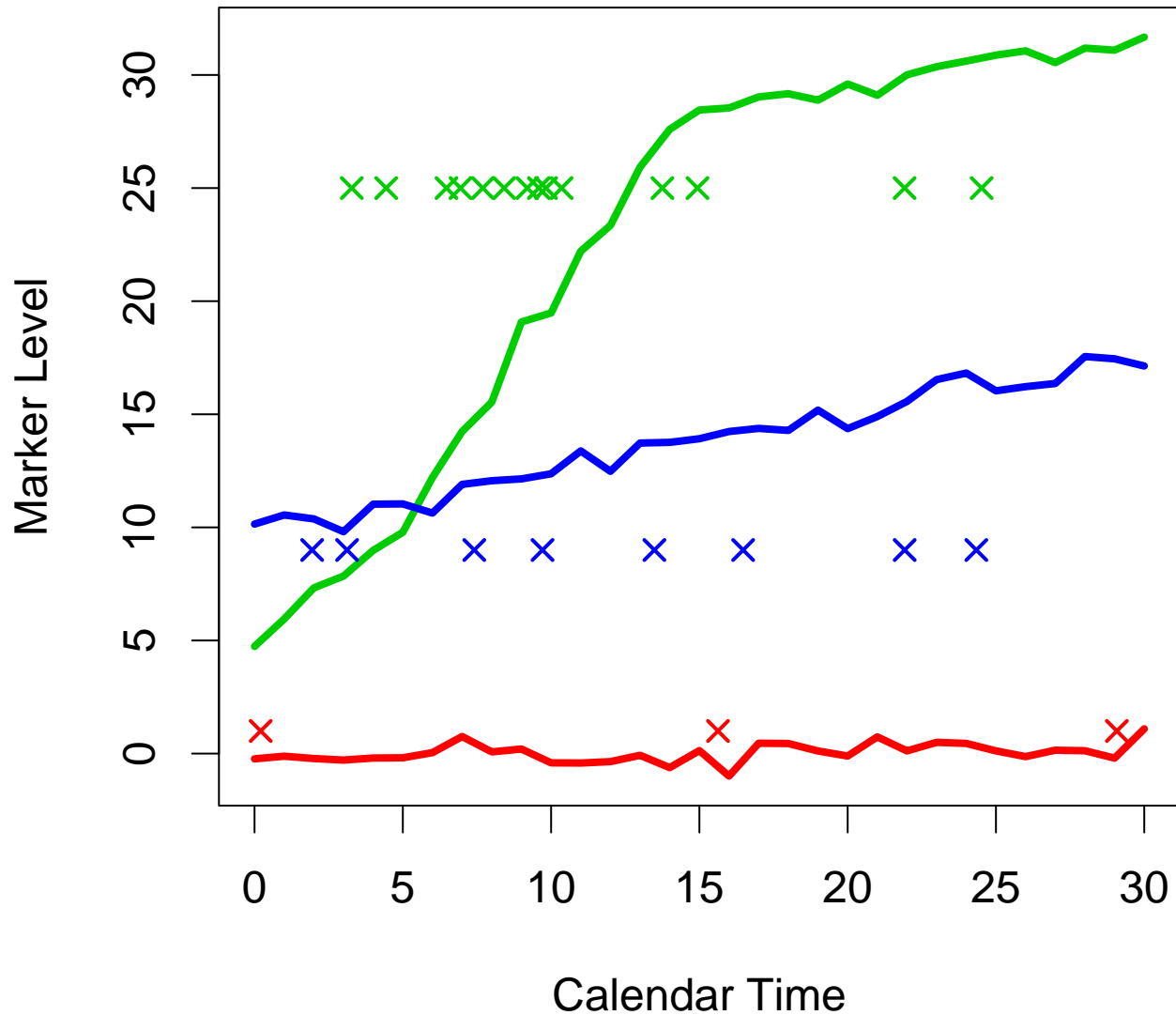
# Advantages of latent classes

- Conditional independence simplifies modeling the association.
- Association can be distinct for each class – “protects” against potential misspecification in other approaches.
- Captures subpopulation structure – distinct patterns for the joint responses.

# Flexibility of latent class approach



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# Single event model: PSA–PCa Lin et al., JASA, 2003

$Y_i = \ln(\text{PSA}_i + 1)$ ,  $T_i = \text{PCa diagnosis time}$

$$[Y_i, T_i] = \sum_{k=1}^K [Y_i | \mathbf{c}_{ik}] [T_i | \mathbf{c}_{ik}] [\mathbf{c}_{ik}]$$

**Latent class membership:** Multinomial logit with covariates  $\nu_i$

$$\Pr(c_{ik} = 1) \propto \exp(\nu_i^T \eta_k)$$

**Longitudinal marker:** Linear with subject- and class-specific effects; covariates  $U_i, V_i, W_i$

$$E(Y_i | b_i, c_{ik} = 1) = U_i \beta + V_i b_i + W_i M_k + \epsilon_i$$

**Hazard:** Cox model with class-specific baseline, frailty ( $Z_i$ ); covariates  $X_i$

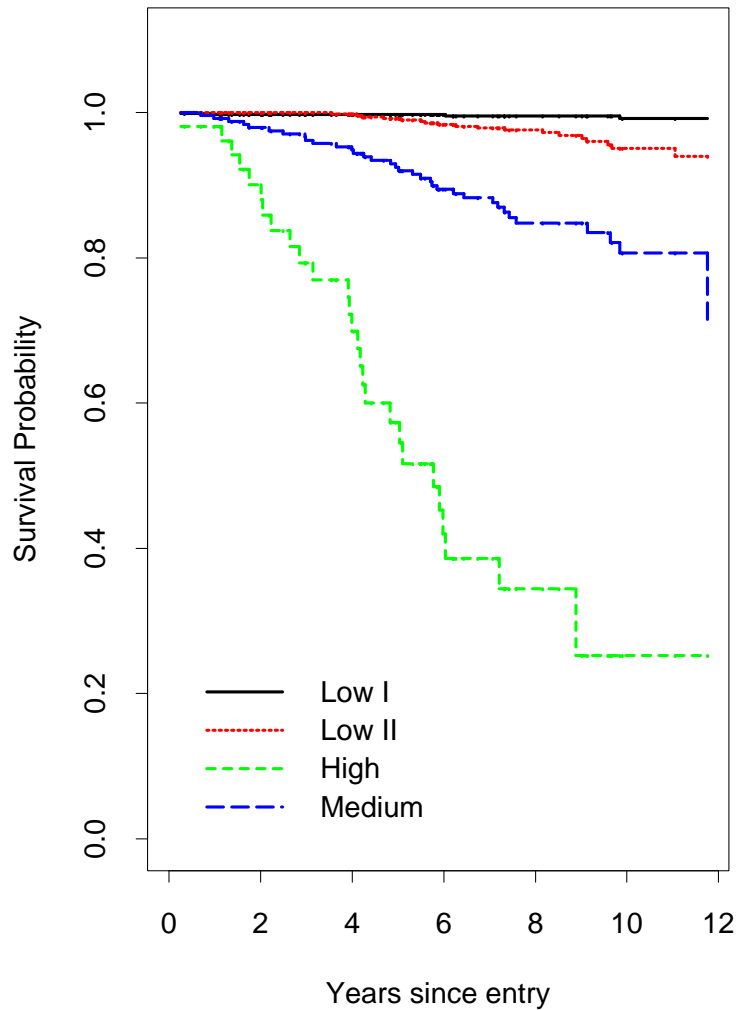
$$Z_i R_i(t) \lambda_k(t) \exp(\gamma^T X_i)$$

# Estimation

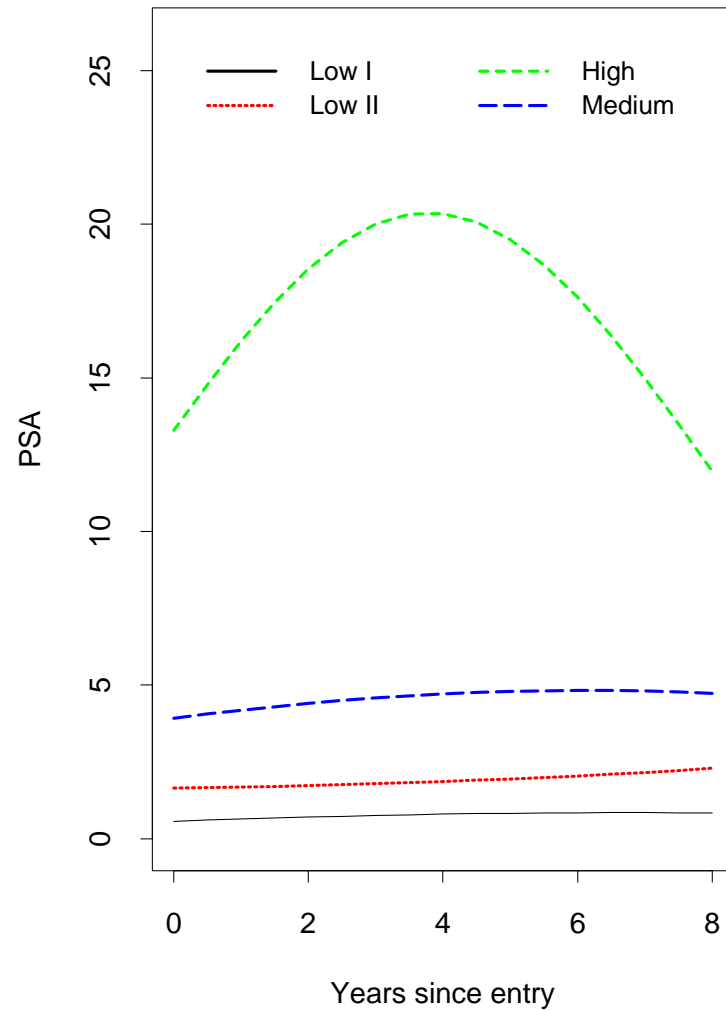
- For fixed number of latent classes  $K$ , use EM algorithm.
  - Complete data:  $(\mathbf{Y}_i, \min\{T_i, \tau_i\}, \delta_i, b_i, Z_i, \mathbf{c}_i)$
  - Missing data:  $(b_i, Z_i, \mathbf{c}_i)$
- Select number of latent classes!  
We used  $BIC = \log L_k - \frac{1}{2}(Npar_k) \log n$

# PSA-PCa results: $K = 4$

Fitted Survival Curves Across 4 -Classes



Fitted Trajectories for the 4 -Class Model



# Proposed Extension

In the latent class structure, replace the single event survival model by the general class of recurrent event models of Peña and Hollander (2003).

$$[\mathbf{Y}_i, (N_i, R_i)] = \sum_{k=1}^K [\mathbf{Y}_i \mid \mathbf{c}_{ik}] [(N_i, R_i) \mid \mathbf{c}_{ik}] [\mathbf{c}_{ik}]$$

- Latent class membership:

$$\Pr(c_{ik} = 1) \propto \exp(\nu_i^T \eta_k)$$

- Longitudinal marker:

$$E(\mathbf{Y}_i \mid b_i, c_{ik} = 1) = U_i \beta + V_i b_i + W_i M_k + \epsilon_i$$

- Recurrent event process:

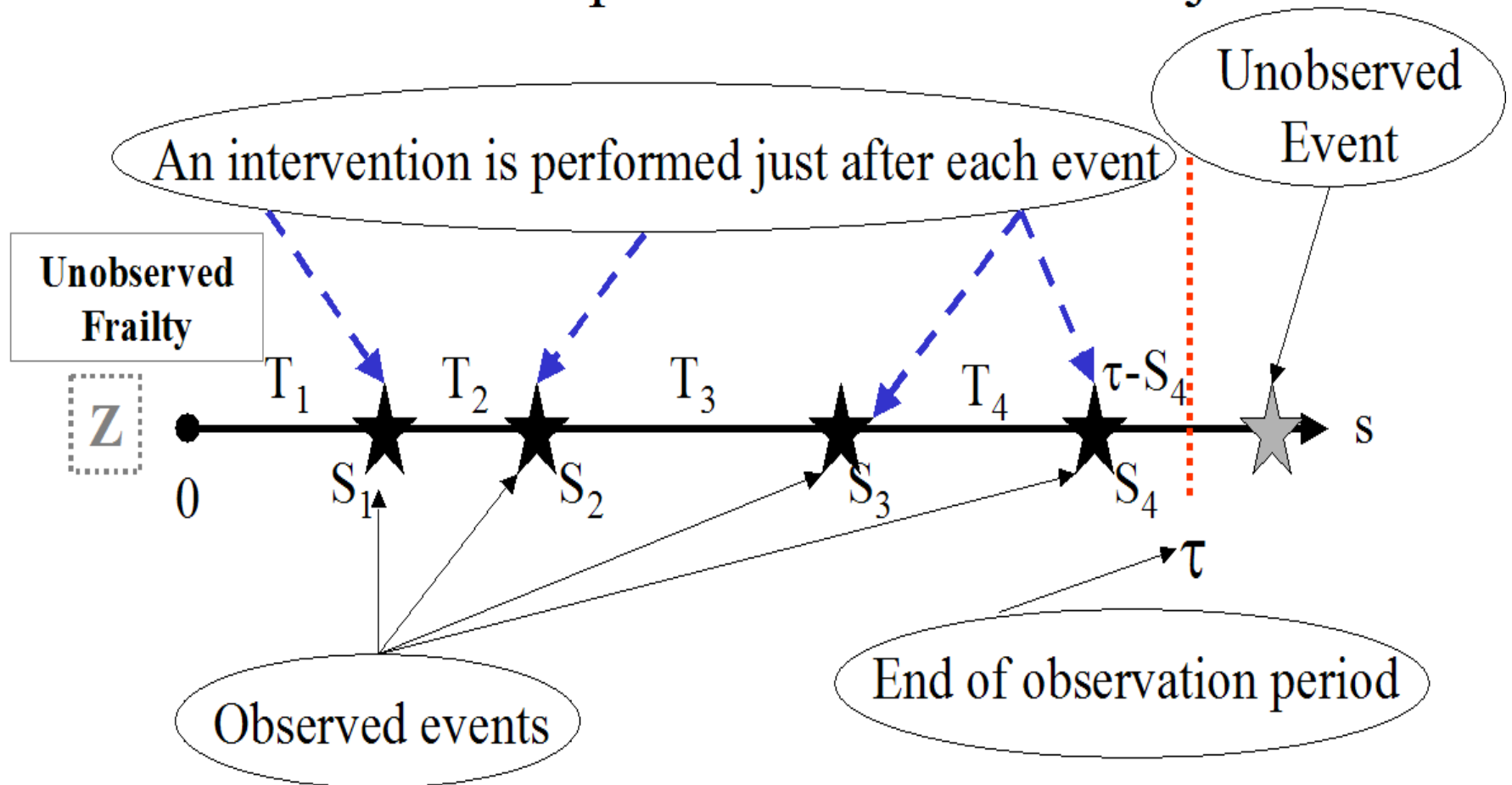
Peña and Hollander (2003)...

# Recurrent event model Peña & Hollander, 2003

Accommodate:

- Effects of intervention following event occurrence.
- Effects of accumulating event occurrences (weakening/strengthening).
- Covariates.
- Random observation time.
- Informativeness of number of events observed.

# A Pictorial Representation: One Subject



An observable covariate vector:  $\mathbf{X}(s) = (X_1(s), X_2(s), \dots, X_q(s))^t$

# Intensity Process (details Saturday)

For subject  $i$  in latent class  $k$ :

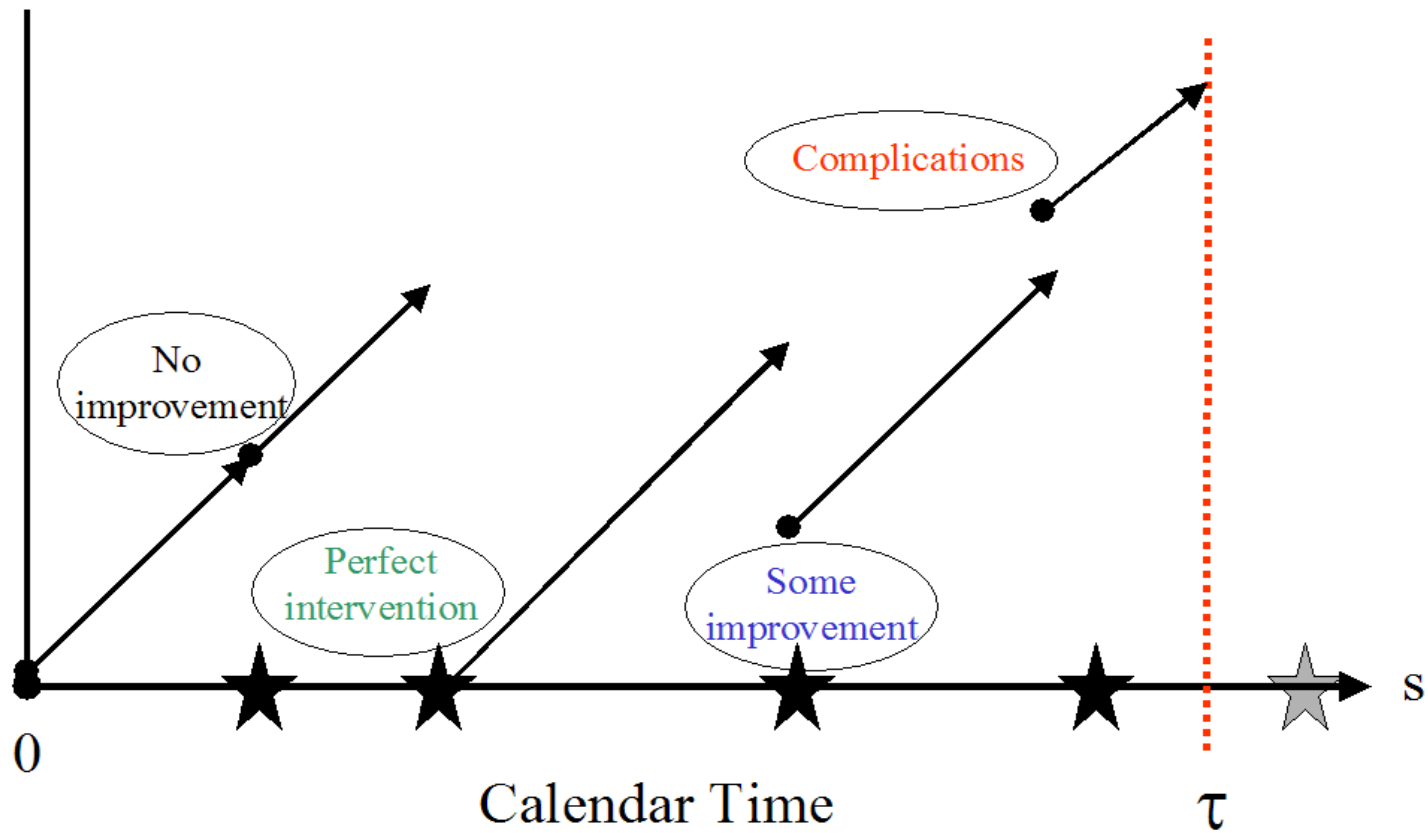
$$\lambda(s|Z_i, c_{ik} = 1) = Z_i \lambda_{0k}[\mathcal{E}_i(s)] \rho[N_i(s-); \alpha_k] \psi[X_i(s)\gamma_k]$$

where

- $\lambda_{0k}(\cdot)$  = unknown baseline hazard rate function for class  $k$ .
- $\mathcal{E}_i(s)$  = **effective age** of the subject at calendar time  $s$ . **Idea:** performed intervention changes the effective age of the subject acting on the baseline hazard rate.
- $\rho(\cdot; \alpha_k)$  = a positive function on  $\{0, 1, 2, \dots\}$  with  $\rho(0; \alpha_k) = 1$  and with unknown parameter  $\alpha_k$ , depending on the class.
- $\psi(\cdot)$  = positive link function containing the effect of covariates  $X_i$ ;  $\gamma_k$  are unknown. e.g.,  $\psi(X_i \gamma_k) = \exp(X_i \gamma_k)$ .
- $Z_i$  = unobservable frailty. Induces associations among subject's inter-event times. Class dependent...?

# Illustration: Effective Age Process “Possible Intervention Effects”

Effective Age,  $E(s)$

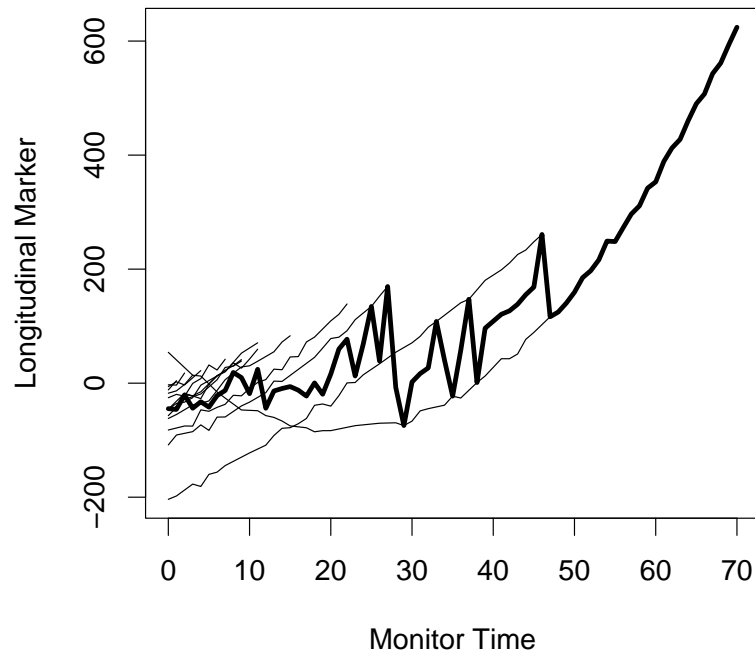


# Illustration

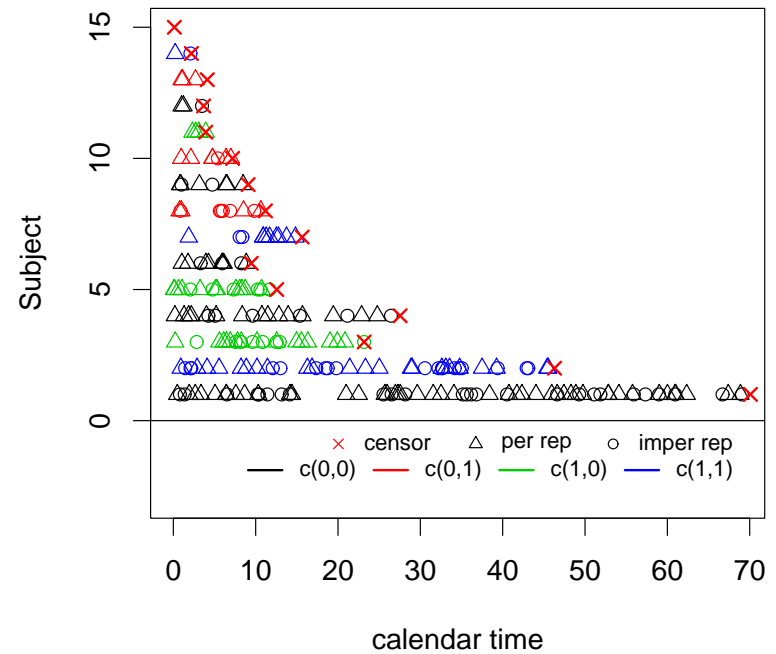
- $K = 3$  latent classes. Class membership based on bivariate binary covariate  $\nu_i$
- Longitudinal marker covariates:
  - Fixed effects:  $U_i = \text{intercept, time, time}^2$ , binary  $\nu_i$
  - Random effects:  $V_i = \text{intercept, time, time}^2$
  - Class effects:  $W_i = \text{intercept, time, time}^2$
- Recurrent event process:
  - $\lambda_{0k}$ : Weibull, params  $(1, 1)$ ,  $(2, 1)$ ,  $(0.5, 1)$
  - $\rho(m) = \alpha_k^m$ ,  $\alpha_1 = 1$ ,  $\alpha_2 = 0.95$ ,  $\alpha_3 = 1.02$
  - $\psi(X_i \gamma_k) = \exp(X_i \gamma_k)$ ,  $X_i = \nu_i$ ,  $\gamma_k \in \mathcal{R}^2$  class dependent
  - $\mathcal{E}(\cdot)$ : Pr(perfect repair) = 0.6
  - No frailty

# Illustration – Class 1

Class = 1; Fixed =  $c(0, 10, 0.2, 1, 0)$ ;  
Class =  $c(0, 0, 0)$

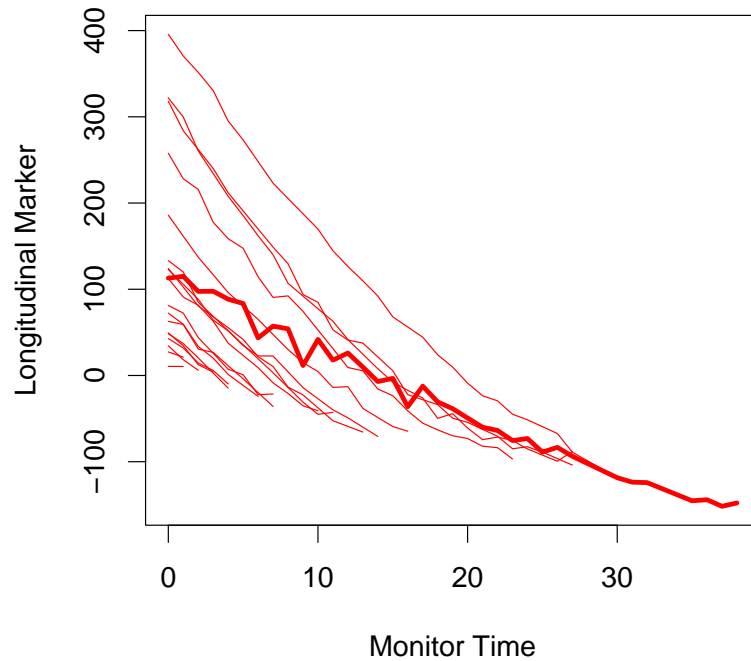


Class = 1; Type = WEI; BaseHaz param =  $c(1, 1)$   
Psi param =  $c(0, 0)$ ; Rho = 1

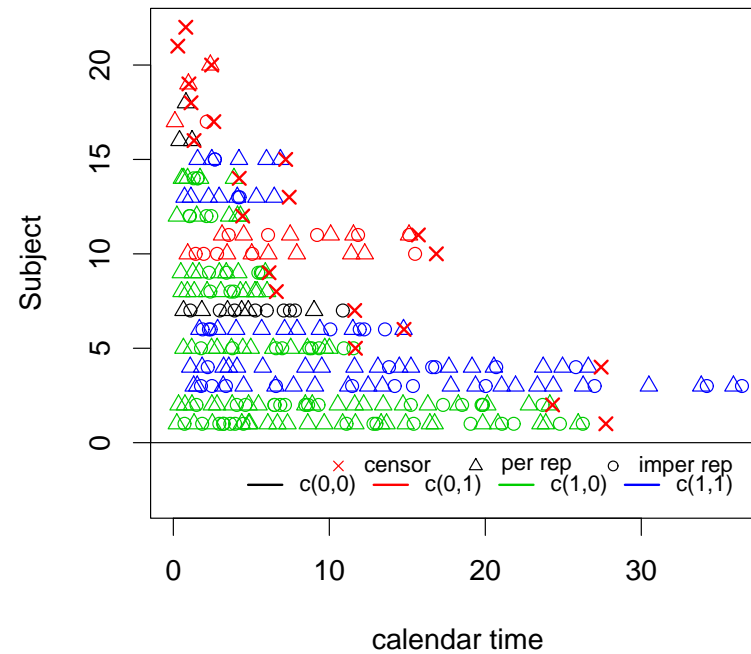


# Illustration – Class 2

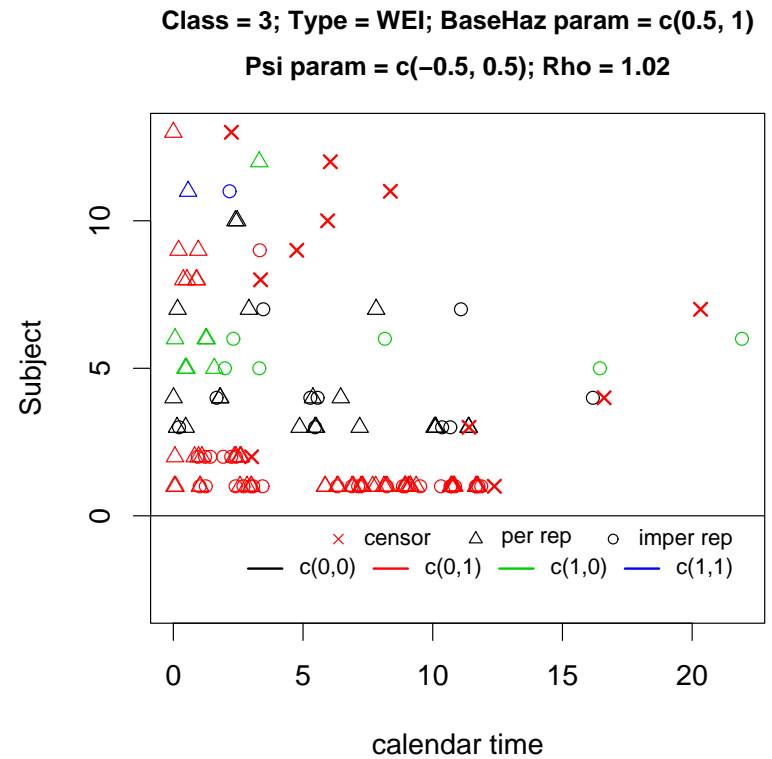
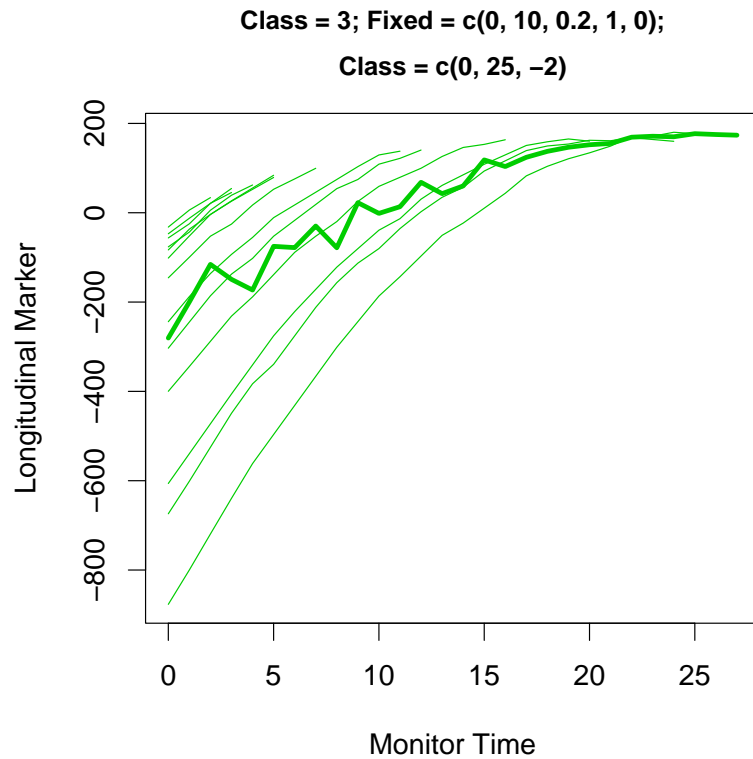
Class = 2; Fixed =  $c(0, 10, 0.2, 1, 0)$ ;  
 Class =  $c(20, -25, 0.2)$



Class = 2; Type = WEI; BaseHaz param =  $c(2, 1)$   
 Psi param =  $c(1, -1)$ ; Rho = 0.95



# Illustration – Class 3



# Estimation

- Maximum likelihood – EM algorithm (Dempster, Laird and Rubin, 1977)
  - Complete data:  
 $(\mathbf{Y}_i, \{(N_i(s), R_i(s)), s \leq \tau_i\}, b_i, Z_i, \mathbf{c}_i)$
  - Missing data:  $(b_i, Z_i, \mathbf{c}_i)$
- Number of latent classes.
  - AIC, BIC
  - Bayesian model averaging
- In progress...

# Diagnostics

- Conditional independence of  $Y_i$  and  $\{(N_i(s), R_i(s)), s \leq \tau_i\}$   
Build upon ideas of generating pseudo class membership
  - Bandeen-Roche *et al.* (1997)
  - Lin *et al.* (2002)
- Residual diagnostics
  - Survival-type residuals *e.g.* (Lin *et al.*, 2002)
  - Build on current work in growth mixture modeling:  
B. Muthén, C. Hendricks Brown, C-P. Wang

# Extensions

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Many!

- Effect of event occurrence (and subsequent intervention) on longitudinal marker.
- Multiple markers (Lin *et al.* 2002).
- Partially observable classes (cure-rate models) (Law *et al.*, 2002; Ibrahim *et al.*)